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# Model-free inference of memory in conformational dynamics of a multi-domain protein

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#### Abstract

Single-molecule experiments provide insight into the motion (conformational dynamics) of individual protein molecules. Usually, a well-defined but coarsegrained intramolecular coordinate is measured and subsequently analysed with the help of hidden Markov models to deduce the kinetics of protein conformational changes. Such approaches rely on the assumption that the microscopic dynamics of the protein evolve according to a Markov-jump process on some network. However, the manifestation and extent of memory in the dynamics of the observable strongly depends on the chosen underlying Markov model, which is generally not known and therefore can lead to misinterpretations. Here, we combine extensive single-molecule plasmon ruler experiments on the heat shock protein Hsp90, computer simulations, and theory to infer and

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quantify memory in a model-free fashion. Our analysis is based on the bare definition of non-Markovian behaviour and does not require any underlying model. In the case of Hsp90 probed by a plasmon ruler, the Markov assumption is found to be clearly and conclusively violated on timescales up to roughly 50 s, which corresponds roughly to  $\sim$ 50% of the inferred correlation time of the signal. The extent of memory is striking and reaches biologically relevant timescales. This implies that memory effects penetrate even the slowest observed motions. We provide clear and reproducible guidelines on how to test for the presence and duration of memory in experimental single-molecule data.

Keywords: memory, protein dynamics, non-Markovian behaviour, plasmon ruler

#### 1. Introduction

Dynamics of biological systems, including data obtained in single-molecule experiments, are often analysed by (Hidden) Markov modelling [1, 2]. This method allows to infer mesoscopic states and Markovian transition rates between them from time series, which can be obtained from time-resolved distance measurements, e.g. by Förster resonance energy transfer (FRET) [1, 2], plasmon rulers [3], and optical [4, 5] or magnetic tweezers [6]. Markov models have been used extensively for analysing such data but their validity critically hinges on the assumption that the observed process is memoryless, meaning that transitioning from one state to another is independent of previous states and that the dwell times within states are exponential random variables [7]. The latter assumption in particular also implies that the transition paths between meso-states are infinitely fast [7]. If either (or both) of the above assumption is violated, the interpretation in terms of Markov models may lead to inconsistent conclusions (see e.g. [8]). Parametrically, memory effects in *observed* dynamics can be incorporated by hidden Markov models (HMMs). However, the magnitude and duration of memory in the observed dynamics strongly depend on the choice of the underlying (microscopic) Markov model. Importantly, the magnitude and extent of memory in observed dynamics has not been consistently tested in any of these experimental studies. As a result, the sensibility and accuracy of HMMs of single-molecule observations [1, 2] has not yet been scrutinised.

Generally, memory is defined as the absence of Markov behaviour, and can emerge from 'time lumping' or 'state lumping' [7-13]. Microscopic classical as well as quantummechanical dynamics is deterministic and thus nominally Markovian. Memory effects in physical systems are thus strictly the results of some sort of lumping of microscopic states [11, 12, 14, 15]. Time lumping directly results from biophysical measurements as the time resolution is not infinite and the signal is binned over small time windows. State lumping, on the other hand, refers to the states of the observed system, and also arises as a consequence of the measurement procedure: if an intramolecular distance of a biomolecule, for example, is observed over time by single-molecule FRET (smFRET) or plasmon ruler, the experimental information refers to only one spatial coordinate, which is the 'projection' of many underlying conformational states. Memory can emerge when two prerequisites are met: first, when there are more degrees of freedom than those observed, signifying that an observed state is a projection of several coordinates onto e.g. only one. Second, when dynamically distinct microscopic states become projected onto the same observable state (for a specific *minimal* example on how memory may emerge upon projection of 'microscopic' Markov-jump dynamics see figure 3 as well as [16]). This holds true in the aforementioned example of the biomolecule as it consists of numerous atoms (i.e. it has numerous degrees of freedom) and exhibits (well-defined) longlived conformational meso-states, but only the distance between two amino acids is observed. In this situation, memory is automatically induced unless all hidden degrees of freedom relax infinitely fast, such that the observable 'feels their presence only on average'.

One important question emerging from this consequential relation is whether such memory effects are biologically relevant. Therefore, it is crucial to quantify both their magnitude and duration. If they relax much faster than the observed biological processes, they can be safely neglected. Conversely, if they relax on a slower timescale, then they may directly affect biological processes as, e.g. protein folding, catalysis or signalling where the relevant timescale is often set by some barrier crossing event [17]. In practice, this is hardly ever known *a priori*, and neither the exact microscopic dynamics nor their projection (i.e. the mapping between full and projected dynamics) are experimentally accessible. This in turn poses the challenging question *how to detect and quantify memory without knowing the underlying microscopic dynamics*?

Non-exponential dwell time distributions in given meso-states in principle (i.e. if present) serve as an indicator for the presence of memory. However, an exponential dwell (i.e. exit time) distribution in turn is *not* an indicator for Markovian dynamics, as memory can also emerge in the sequence of consecutively visited states (see example in [8]).

A preliminary upper bound on the duration of memory can be inferred from the correlation (or 'mixing') time  $t_{corr}$  of the signal. The (auto)correlation time may be defined via the (auto)correlation function  $C(t) = \langle q_t q_0 \rangle - \langle q_t \rangle \langle q_0 \rangle$ , which is determined as the average over a (large) ensemble of N observed/projected trajectories or the time average over a sufficiently long trajectory, i.e.

$$C(t) = \lim_{N \to \infty} \left( N^{-1} \sum_{i=1}^{N} q_t^i q_0^i - N^{-2} \sum_{i=1}^{N} q_t^i \sum_{j=1}^{N} q_0^j \right)$$
  
= 
$$\lim_{T \to \infty; T/t \gg 1} (T-t)^{-1} \int_0^{T-t} q_{\tau+t} q_\tau d\tau - \left( \lim_{T \to \infty} (T)^{-1} \int_0^T q_\tau d\tau \right)^2, \quad (1)$$

where we assume that the observed process  $q_t$  is ergodic and that the *microscopic* dynamics are stationary and time-homogeneous. The autocorrelation C(t) of a time-series measures the correlation between the coordinate q at t = 0 and different time points t. In essence, it quantifies the 'degree of similarity' between these lag times. As such, it captures the extend of 'memory' of both the observable and hidden initial conditions. The correlation time is defined as (see e.g. [18])

$$\frac{1}{t_{\rm corr}} \equiv -\lim_{t \to \infty} \frac{1}{t} \ln C(t)$$
<sup>(2)</sup>

and reflects the timescale on which the entire system (i.e. observable plus any hidden degrees of freedom) de-correlates from the initial condition. Alternatively, one can instead define a normalised autocorrelation function  $C(t) \equiv C(t)/C(0)$ . Such an analysis is very helpful to obtain a first upper bound on the memory timescale, since (by definition) if memory exists the correlation time is the slowest timescale in the system. However, it is important to note that the correlation time does not equal the memory time (see e.g. [19]). In fact, it does not even imply the presence of memory. To see this, recall that a one-dimensional diffusion in a bi-stable or even in a single-well parabolic potential both have a non-vanishing correlation time while they are perfectly Markovian. Moreover, conclusions drawn form so-called 'normalised correlation times' [20, 21],  $\hat{t}_{corr} \equiv \int_0^\infty C(\tau) d\tau$ , may lead to misconceptions about the underlying dynamics when the latter span multiple timescales [19]. A correlation time analysis is also not suitable for extracting the temporal extent of memory when the latter is present, because it is only an upper bound and therefore has no implications on the duration of memory.

An elegant indicator of memory in low-dimensional reaction coordinates focuses on the fluctuations of transition-path times [22] (see also generalisation for dynamics on general graphs in [7]). This analysis exploits the fact that transition-path times—the durations of successful transitions—for a one-dimensional Markov process cannot have a coefficient of variation exceeding one. Thus, values exceeding one are a reliable indicator of memory. However, a coefficient smaller than one does not necessarily imply Markov behaviour and this analysis also requires superb temporal resolution (much faster than the transition duration), which is often not feasible experimentally.

A more robust (concerning temporal resolution) and model-free (but still not ideal; see below) test for the presence and duration of memory is to test *directly* for the violation of Markov behaviour [18]. This method offers significant flexibility and interpretability, free from the model-selection challenges [23] inherent to model-based approaches like HMMs. The approach rests on one general consequence of the Markov property encoded in the socalled Chapman–Kolmogorov equation. That is, the (two-point) transition probability density  $G(q,t|q_0,0)$  for a transition  $(q_0,0) \rightarrow (q,t)$  of any Markov process satisfies

$$G(q,t|q_{0},0) \stackrel{\text{Markov}}{=} \int dq' G(q,t|q',\tau) G(q',\tau|q_{0},0)$$

$$\stackrel{t-\text{hom}}{=} \int dq' G(q,t-\tau|q',0) G(q',\tau|q_{0},0)$$
(3)

where *t*-hom denotes time-homogeneous Markov dynamics. Note in particular that there is no dependence on  $\tau$ . Since the Chapman–Kolmogorov equation holds true for all Markov processes, one can turn this property into a test for Markovianity by determining the transition probability density  $G(q,t|q_0,0)$  for the actual *observed* (generally non-Markovian) dynamics, and compare it with the Chapman–Kolmogorov construct

$$G_{\tau}^{\rm CK}(q,t|q_0) \equiv \int \mathrm{d}q' G(q,t-\tau|q',0) \, G(q',\tau|q_0,0) \,. \tag{4}$$

The Chapman–Kolmogorov construct  $G_{\tau}^{CK}$  corresponds to a fictitious process in which we force all hidden degrees of freedom to their stationary distribution at time  $\tau$ , and thereby erase all memory (if present). The transition probability density can be determined from the recorded (experimental or simulated) time series  $\{q_t\}_{0 \le t \le T}$  (one-dimensional in our case) of total duration T = 6h (e.g. see figures 2(d)–(f) or figure 4(a). Assuming that the time-series is stationary, which is our only assumption that we also verified to be the case here by confirming that waiting times in the states do *not* explicitly depend on time, we can determine the two-point transition probability density as

$$G(q,t|q_0,0) = \frac{\langle \delta(q_{t+\tau} - q) \,\delta(q_{\tau} - q_0) \rangle}{\langle \delta(q_{\tau} - q_0) \rangle},\tag{5}$$

where as before  $\langle \cdot \rangle$  denotes an ensemble or time average over a long trajectory and  $\delta(x)$  is Dirac's delta distribution. We note that in the analysis we lump the data into bins as we are limited by the sampling. Generally, the binning itself may introduce memory [13, 24], which can be gauged by repeating the analysis for different bin sizes (see below). However, the analysis itself rests on no such assumption, i.e. it can be performed on an arbitrarily fine binning as long as sampling permits a reliable inference of  $G(q,t|q_0,0)$ . Due to our assumption that  $q_t$  is ergodic, the system reaches a stationary equilibrium density  $p_{eq}(q)$  for ergodically long times, i.e.  $\lim_{t\to\infty} G(q,t|q_0,0) = p_{eq}(q)$ . Note that, given sufficient statistics, one can infer  $G(q,t|q_0,0)$  and  $G_{\tau}^{CK}(q,t|q_0)$  directly from the measured time series  $q_t$ , and that in particular, there is no need to specify any underlying model for the observed (projected) and microscopic dynamics. The comparison between G and  $G_{\tau}^{CK}$  may be done via the Kullback–Leibler divergence (also called relative entropy) [25]

$$\mathcal{D}_{\tau,q_0}(t) \equiv \int \mathrm{d}q G(q,t|q_0,0) \ln\left[\frac{G(q,t|q_0,0)}{G_{\tau}^{\mathrm{CK}}(q,t|q_0)}\right] \ge 0.$$
(6)

By construction,  $\mathcal{D}_{\tau,q_0}(t) = 0$  if and only if  $G(q,t|q_0) = G_{\tau}^{CK}(q,t|q_0)$  for all q. That is,  $\mathcal{D}_{\tau,q_0}(t)$  will by construction vanish for a Markov process. A non-zero value of  $\mathcal{D}_{\tau,q_0}(t)$  for any  $\tau$  and t thus reflects genuine memory in the dynamics in the sense that as soon as  $\mathcal{D}_{\tau,q_0}(t) \neq 0$  for some  $\tau$  and  $q_0$  the observed process  $q_t$  is conclusively non-Markovian. Note, however, that there are non-Markovian processes which also satisfy the Chapman-Kolmogorov equation (3) [26–28]—and thus would also result in a vanishing  $\mathcal{D}_{\tau,q_0}$ . Therefore, while  $\mathcal{D}_{\tau,q_0}(t) \neq 0$  implies memory, the converse  $\mathcal{D}_{\tau,q_0}(t) = 0$  does *not* necessarily imply that the observed process is Markovian.

Notably, for long times  $t \ge t_{corr}$ , and similarly for  $\tau \ge t_{corr}$  (and  $t = \tau$  by construction), one expects that  $\mathcal{D}_{\tau,q_0}(t)$  approaches zero. Analysing  $\mathcal{D}_{\tau,q_0}(t)$  for different  $\tau$  and  $q_0$  thus allows us to quantify the magnitude and temporal extent of memory in a model-free manner (see below). The Kullback–Leibler divergence in equation (6) reflects how much the observable  $q_t$  at a given time t 'remembers' the initial state of *hidden* degrees of freedom and quantifies how memory attenuates in the course of time while microscopic trajectories gradually mix in configuration space. Lastly, it is important to highlight that the information encoded in the temporal evolution of  $\mathcal{C}(t)$  and  $\mathcal{D}_{\tau,q_0}(t)$  is fundamentally different. This distinction becomes evident in figures 2 and 3, which compare a Markovian with non-Markovian observations. Even in the absence of memory, there exists a timescale on which the coordinate  $q_t$  forgets the initial conditions. However, in the Markovian case, there is no memory and  $\mathcal{D}_{\tau,q_0}(t) = 0$  as required. Therefore,  $\mathcal{D}_{\tau,q_0}(t)$  in contrast to  $\mathcal{C}(t)$ , indeed provides information about the correlation between the observable  $q_t$  and the initial conditions of *hidden* degrees of freedom, thereby offering genuine information about the extent and duration of memory.

Here, we investigate memory effects in single-molecule time series of the conformational dynamics of the protein Hsp90 probed by a plasmon ruler, monitoring the distance between a pair of plasmonic nanoparticles attached to the protein (see figure 1), as well as a set of HMMs of this observable. Hsp90 is a molecular chaperone (a so-called 'helper protein'), as such, it helps other proteins to fold properly and to achieve an active (i.e. biologically functional) conformation. As a heat shock protein, its production is modified by stress conditions. Hsp90 is highly conserved and one of the most abundant proteins within the cytoplasm of cells [29–31]. Renowned for its extensive interaction network with numerous cochaperones (one may call them 'helper's helpers') and protein clients (also called substrates), Hsp90 plays a pivotal role in the maintenance of essential cellular functions [32, 33]. Structurally, Hsp90 exists as a homodimer, with each monomer composed of three domains, as shown in figure 1: the N-terminal domain (NTD) containing an ATP-binding site, the middle domain (MD) crucial for ATP-hydrolysis and linked to the NTD via a charged linker sequence, and the Cterminal domain (CTD), serving as the dimerisation interface with the highest affinity [34, 35]. The Hsp90 dimer can undergo large conformational changes, transitioning from a V-shaped N-terminally open structure to a tightly closed state, as evidenced by a wealth of structural



**Figure 1.** The heat shock protein Hsp90 (green-blue) is a homodimer composed of two identical proteins, which each have a N, M and C-domain. They are bound to each other at their C-termini with a high affinity coiled-coil linker. The protein exhibits N-terminal opening and closing dynamics, as illustrated by the Hsp90 cartoons on the right. This motion of a single dimer can be recorded by plasmon ruler spectroscopy in a dark field microscope where two gold nanoparticles are attached to the protein. As the coupling strength of the gold particles depends on their distance, a smaller distance results in stronger coupling and therefore larger scattering intensity (plasmon signal (q)).

data [36, 37]. Several investigations have measured time series of N-terminal opening and closing with single-molecule FRET [38–41]. These studies are limited by photobleaching, resulting in insufficient statistics to assess the Markov assumption. However, plasmon ruler spectroscopy allows for a much longer observation of Hsp90's conformational changes, monitoring its opening and closing dynamics in the range of hours [3]. In the following, we analyse some of the plasmon ruler data from [3]. The plasmon data is available at [42].

Signatures of memory in the time-series of Hsp90 were already observed in [3] in terms of non-exponential dwell times in the two respective conformational meso-states. Non-Markovian behaviour is indeed anticipated to manifest since many protein conformations lead to the same observable plasmon signal  $q_t$ , here a proxy for the distance between plasmonic particles, while microscopic trajectories evolving from an ensemble of conformations with the same distance between plasmonic particles will be dynamically distinct. The magnitude and timescale on which memory emerges remain unknown.

#### 2. Simulated single-molecule time series

#### 2.1. Data simulation

In order to examine the consistency and limitations of this approach, which requires lots of data, we start with simulated time-series of model systems both with and without memory, and in or driven out of equilibrium. All the three models consist of four states each (figures 2(a)–(c)) and are described by Markovian jump dynamics on the microscopic level. Moreover, we consider dynamics on a discrete time, 'probed' with a frequency of 10 Hz, since in the practical context of biophysical measurements (see section 3) the recordings are inherently discrete (which here in the limit of long measurements is expected to essentially become equivalent to continuous-time dynamics, e.g. Gillespie simulations [43, 44]). The simulation of time-series for the full dynamics was done with MASH-FRET [45], which includes (Poissonian) photon



**Figure 2.** Examples of simulated time-series for three different four-state models projected onto two observable states. Each time-series  $\{q_i\}_{0 \le t \le T}$  of total duration T = 6 h (only 10000s depicted) is shown for a four-state Markov model with (a+d) uniform transition rates, (b+e) with Hsp90-inspired transition dynamics that obey detailed balance, (c+f) with Hsp90-inspired transition dynamics that is driven out of equilibrium. (g) Probability density functions  $p_{eq}(q)$  of one example time-series for each of the introduced model systems (almost lying on top of each other). The two peaks correspond to the closed (higher q) and open (lower q) conformation. (h) Averaged autocorrelation function C(t) of simulated time-series of the corresponding models. The timescale  $t_{corr}$  is identified from the long-time behaviour of C(t) as detailed in the main text.

shot noise, similar to the noise occurring in single-molecule FRET time series, yielding a signal that is *not* dichotomous anymore.

The respective model parameters (in particular signal intensities) were chosen to mimic the distinct protein conformations observable in the experiments. Corresponding kinetic transitions rates between states were either chosen to be identical (figure 2(a)), or inspired by those recently obtained from single-molecule FRET data of Hsp90's conformational dynamics (figures 2(b) and (c)) [41]. In the latter case, the first system (see figure 2(b) models Hsp90 dynamics in equilibrium (i.e. transition rates obey detailed balance) whereas the second system (see figure 2(c) is out of equilibrium with an cycle-affinity of 10  $k_BT$  [46].

To allow for a better comparison with experimental data in later sections, all simulated 4-state models were divided into two low intensity (centred at q = 0.05; open) and two high

intensity (centred at q = 0.95; closed) states which are consecutively lumped together to create hidden states, giving rise to only two observable meso-states. The resulting 'projected' dynamics  $q_t$  (figures 2(d)–(f)) thus now mimic the effective two-state dynamics observed in the experimental setup (compare figure 4(a). Recall that the noise introduced by MASH-FRET blurs the discrete-state trajectory, ultimately giving meso-states a finite extension as seen in the stationary histograms  $p_{eq}(q)$  in figure 2(g).

The following memory analysis will be carried out for data sets comprising 100 independent time-series for each system with a duration of T = 6 h and time resolution of 10Hz each—identical parameters as in the experimental case. Simulation and analysis of the data were performed independently by different researchers. The person analysing the data thus had no prior knowledge of the underlying system.

#### 2.2. Analysis of simulated single-molecule time series

We start our analysis by first determining the normalised autocorrelation function C(t) (see equation (1)) for each individual time-series of the three introduced models, and consecutively take the corresponding sample-average over the respective 100 independent measurements (see figure 2(h)). Furthermore, the correlation times  $t_{corr}$  are determined by fitting the long-time behaviour of C(t) to an exponential  $C(t) \simeq \exp(-t/t_{corr})$ . The inferred values are  $t_{corr} \approx 49.1$  s,  $t_{corr} \approx 14.5$  s, and  $t_{corr} \approx 3.4$  s, respectively, as indicated by the markers in figure 2(h) (further details are given in table 1).

Next we quantified the memory via the Chapman-Kolmogorov construction 4 (see figure 3). The transition probability density  $G(q,t|q_0,0)$  for the recorded time-series was determined according to equation (5) via a histogram analysis. Since the probability density functions  $p_{eq}(q)$  of all considered models (also for the experimental data; see section 3) are bimodal, indicating two well-defined and well-separated meso-states (see figure 1), we start by lumping the state-space into two equal bins of width  $l_{\text{bin}} = 0.5$ , centred at 0.25 and 0.75 respectively. As a result, q is now discrete-valued,  $q = \{0, 1\}$ , and refers to a pair of bins at 0.25 and 0.75 with a width  $l_{\rm bin} = 0.5$ . Each bin encompasses one of the two wells, such that  $q_0 = 0.25$  corresponds to the open state and  $q_0 = 0.75$  to the closed state throughout, respectively. Note that the precise position of q is thus immaterial. Consequently, the reference Chapman–Kolmogorov construction and in turn also the Kullback–Leibler divergence  $\mathcal{D}_{\tau,q_0}(t)$  are straightforwardly obtained via equations (4) and (6) by replacing integrals by sums over the two lumps. For the following analysis, we choose  $q_0 = 0.25$  and for completeness show the corresponding results for  $q_0 = 0.75$  in the appendix since they are almost identical. Moreover, for the experimental Hsp90 time-series (see below) we additionally repeat an analogous memory analysis for a finer spatial lumping into four equal bins of size  $l_{bin} = 0.25$ , and eight equal bins of size  $l_{bin} = 0.125$ , is illustrated in figure 7 of the appendix and displays essentially quantitatively equivalent behaviour. We note that binning alone of truly one-dimensional Markovian dynamics may cause memory effects (see [13]). However, the memory here is genuinely projection-induced (i.e. by observing the plasmon signal only) and is therefore expected to persist even for an 'infinitely fine' binning of the plasmon signal (see e.g. [26]), for which, however, exponentially more data is required for an equivalent analysis.

A priori, the consistency of our approach requires that *no* detection of memory is expected for the first system (figure 2(a)), since in this case all possible microscopic paths are equivalent (i.e. all transition rates are identical). Indeed, as anticipated, the Kullback–Leibler divergence  $D_{\tau,q_0}(t)$  (6) was found to be zero (within statistical errors) for all times *t* and shown values of  $\tau$ (see figure 3 first column). Only for increasingly large values of  $\tau$ , a noisy result was obtained due to undersampling of the functions *G* entering the Chapman–Kolmogorov construction (4)



**Figure 3.** Analysis of memory in projected dynamics for three different simulated model systems with 100 time-series of 6 h duration each. The Kullback–Leibler divergence  $\mathcal{D}_{\tau,q_0}(t)$  in equation (6) between the transition probability density of the observed dynamics  $G(q,t|q_0)$  and its Chapman–Kolmogorov construction  $G_{\tau}^{CK}(q,t|q_0)$  as a function of time *t* for underlying dynamics of (a)–(c) a four-state Markov model with uniform transitions, (d)–(f) a four-state Markov model with Hsp90-inspired transition dynamics that obey detailed balance, and (g)–(i) a four-state Markov model with Hsp90-inspired transition dynamics that is driven out of equilibrium, respectively. Initial condition was set here to  $q_0 = 0.25$ , corresponding to the open conformation of Hsp90 (see appendix for  $q_0 = 0.75$ ). Memory in the Chapman–Kolmogorov construct is reset to zero at different times  $\tau$  and the memory time-scale  $t_{mem}$  (red marker) is obtained by fitting the long-time behaviour according to  $\mathcal{D}_{\tau,q_0}(t) \simeq \exp(-t/t_{mem})$ . Shaded areas depict standard deviations  $\sigma_{\mathcal{D}}(t)$ , obtained by averaging over 100 independent simulated trajectories and insets depict the same analysis for only one trajectory that is 100 times as long. The correlation time  $t_{corr}$  is indicated by a yellow marker.

necessarily yielding a positive  $\mathcal{D}_{\tau,q_0}$ . Note that the uncertainty emerging for large  $\tau$  is due to undersampling, i.e. can be overcome by longer traces (see inset figures 3 and 6). We remark that while no memory was (correctly) identified, the system displayed a correlation time of  $t_{\text{corr}} \approx 49.1 \text{ s}$ , underscoring that the correlation time  $t_{\text{corr}}$  (yellow marker) is a bad proxy for inferring memory. It does provide an upper bound on the duration of memory if the latter is present, but in turn has no implication for the presence of memory.

For the two remaining models with rates inspired by single-molecule FRET data of Hsp90 dynamics in equilibrium (figure 2(b)) and driven out of equilibrium with an affinity of 10  $k_BT$  (figure 2(c)), an analogous analysis confirms the presence of memory. In both instances, the system displays memory identified by a non-vanishing Kullback–Leibler divergence over



**Figure 4.** Overview of experimental single-molecule time-series for the Hsp90 protein. (a) Recorded example time-series  $\{q_t\}_{0 \le t \le T}$  of duration T = 6 h via plasmon ruler spectroscopy [3]. (b) Probability density function  $p_{eq}(q)$  of four recorded Hsp90 time-series  $q_T$ . The two visible peaks correspond to the closed (higher population) and open (lower population) conformation of the Hsp90 protein, respectively. (c) Autocorrelation function C(t) of the recorded experimental time-series. The correlation time  $t_{corr}$  (see equation (2)) is identified via an exponential fit of the long-time behaviour of the (averaged) autocorrelation C(t).

approximately one decade in time (figure 3 2nd and 3rd column). For the system in equilibrium, memory was best detectable for a  $\tau = 1$  s. As before, large values of  $\tau$  gave rise to undersampling and in turn larger fluctuations (compare shaded regions), and the results become noisier and less reliable. When using longer time traces, more data points are available and this problem does not emerge (see insets in figure 4 for single trajectories that are longer by a factor of 100). For practical applications, this means that longer measurements or experiments with a higher time resolution result in a more stable and reliable analysis. In particular, an analysis of long traces performs better than the same amount of data split into several short traces (all longer that the correlation time). This is because for a given lag time *t* in equation (5) one has a total interval T - t available for averaging per trajectory of duration *T*, and thus a total interval  $p \times (T-t)$  for *p* trajectories of length *T* and a total interval  $p \times T - t$  for a single trajectory of length  $p \times T$ . If the correlation time is  $t_{corr}$ , one thus has roughly  $\approx p \times (T-t)/t_{corr}$  statistically independent realisations with a single long trajectory.

In line with the correlation time, one may analogously introduce a memory time-scale  $t_{mem}$  as

$$\frac{1}{t_{\text{mem}}} \equiv -\lim_{t \to \infty} \frac{1}{t} \ln \mathcal{D}_{\tau, q_0}(t), \qquad (7)$$

which is inferred by fitting the longest time-scale to an exponential  $\mathcal{D}_{\tau,q_0}(t) \simeq \exp(-t/t_{\text{mem}})$ . Consistently, we always find  $t_{\text{mem}} \leq t_{\text{corr}}$  where, for the considered parameters of  $\tau$ , the equilibrium system was found to take values in the range of  $t_{\text{mem}} \approx 6.2$  s to  $t_{\text{mem}} \approx 6.5$  s and the driven system comparably  $t_{\text{mem}} \approx 2.1$  s to  $t_{\text{mem}} \approx 6.5$  s (see figure 3 and table 1).

#### 3. Experimental single-molecule time series

#### 3.1. Experimental data acquisition

The experimental data analysed here has been recorded by plasmon ruler spectroscopy [42]. For these measurements, two gold nanoparticles (diameter about 60 nm) were attached to yeast Hsp90 via flexible PEG (polyethylene glycole) polymer linkers, one at each monomer. For the attachment of the PEG, Hsp90 was mutated for a cysteine at position 285 in the M-domain. These constructs were immobilised on a glass substrate and observed in a dark field microscope. Due to plasmon coupling between the two gold nanospheres, their plasmon resonance is shifted to higher wavelengths, and the scattering intensity increases when the interparticle distance is decreased [47]. The scattering intensity is recorded and allows to monitor one spatial coordinate of Hsp90's conformational changes. The open Hsp90 dimer shows a low plasmon scattering intensity in arbitrary units, while the closed dimer gives a high signal (see figure 1). We analysed four time series of duration T = 6h each with a recording rate of 10 Hz from previously published data [3]. All traces were from Hsp90 in equilibrium, i.e. in the absence of nucleotides. One exemplary trace is shown in figure 4(a). The effect of the coupling of the gold nanoparticles to the here investigated Hsp90's dynamics could already be excluded in the original publication [3], with the diffusion time of  $0.1\mu$ s for 60nm gold nanoparticles being around six orders of magnitude faster than the observed Hsp90 dynamics here.

# 3.2. Reliability analysis via bootstrapping

To gauge the reliability and robustness of the results with respect to (an unavoidably) limited sampling, we employ a simple bootstrapping routine. In addition to performing the analysis of the relative entropy in equation (6) inferred from the complete dataset, we repeated the analysis on bootstrapped sets. Precisely, we determined the relative entropy from M = 100 independent bootstrap samples obtained by randomly neglecting 20% of the data. From the respective relative entropies of the bootstrapped samples  $\mathcal{D}_{\tau,q_0}^{(i)}(t)$ , we determine the standard deviation  $\sigma_{\mathcal{D}}$  of  $\mathcal{D}_{\tau,q_0}(t)$  as

$$\sigma_{\mathcal{D}}^{2}(t) \equiv M^{-1} \sum_{i=1}^{M} \mathcal{D}_{\tau,q_{0}}^{(i)}(t)^{2} - \left(M^{-1} \sum_{i=1}^{M} \mathcal{D}_{\tau,q_{0}}^{(i)}(t)\right)^{2}.$$
(8)

#### 3.3. Analysis of experimental single-molecule time series

Figures 4(a) and (b) show a measured trajectory  $q_t$  and the equilibrium probability density functions  $p_{eq}(q)$  of the four analysed trajectories, respectively. As in the simulations, the two peaks correspond to the closed (high q) and open (low q) conformation of the Hsp90 protein. Note that the plasmon signal (here q) was normalised between zero and one and scaled such that the peak positions inferred from individual traces match. The normalised autocorrelation functions C(t) of the four traces shown in figure 4(c) are found to nicely superimpose and display an (averaged) correlation time of  $t_{corr} \approx 90$ s (see table 2), as above, obtained via an exponential fit of the long-time behaviour.

The Kullback–Leibler divergence for different 'memory reset times'  $\tau$  are shown in figure 5 for different initial conditions  $q_0$ , respectively. Note that the memory profile, in particular the magnitude, depends on the initial condition of the observable. That is, when starting in the open state ( $q_0 = 0.25$ , blue curve in figure 5), the maximum of  $\mathcal{D}_{\tau,q_0}(t)$  is smaller than when starting from the closed state ( $q_0 = 0.75$ , orange curve in figure 5). From a biophysical perspective,



**Figure 5.** Analysis of memory in experimental Hsp90 time-series. Kullback–Leibler divergence  $\mathcal{D}_{\tau,q_0}(t)$  in equation (6) between the transition probability density of the observed dynamics  $G(q,t|q_0)$  and its Chapman–Kolmogorov construction  $G_{\tau}^{CK}(q,t|q_0)$  as a function of time *t*. Two different initial conditions  $q_0$  of the Hsp90 protein are considered, corresponding to starting in the left (blue) and right (orange) peak in the probability density  $p_{eq}(q)$ , respectively. Memory in the Chapman–Kolmogorov construction is reset to zero at times (a)  $\tau = 1$  s, (b)  $\tau = 3$  s, and  $\tau = 6$  s, respectively. Error bars depict the standard deviation  $\sigma_{\mathcal{D}}(t)$ , obtained by systematically neglecting 20% of the recorded data for M = 100 independent bootstrapped repetitions and subsequent averaging. Memory time-scales are identified via an exponential fit of the long-time behaviour. Note that  $t_{mem}$  (orange and blue markers) is always smaller than the correlation time  $t_{corr}$  (pink marker).

this is not too surprising, given the structural complexity of Hsp90 and the fact that the open state is much more flexible than the closed state [48]. In any case, the finding seems to indicate that a reduced model for  $q_t$  in terms of memory-kernels must *not* be of convolution type and should depend explicitly on  $q_0$ .

Signatures of memory emerged already at very short times (on the scale of the temporal resolution of the experiment). The maximum was reached after just a few seconds (figure 5). Then, the memory decreased slowly and was completely lost only after around 50s. The memory timescale was again obtained from an exponential fit to the long-term behaviour and was here  $t_{\rm mem} \approx 37.2$  s to  $t_{\rm mem} \approx 47.9$  s, respectively, for different  $\tau$  and  $q_0$  (see figure 5 and table 2). Thus, memory was present for over two decades and most pronounced for the 'well' corresponding to the closed conformation of Hsp90 (orange in figure 5). This could be due to Hsp90 being more flexible in its open state [48] enabling a faster relaxation of hidden degrees of freedom. We stress that the actual memory profile in figure 5 is (qualitatively and quantitatively) very different from the one predicted by the Hsp90-inspired HMMs in figure 3. In particular, the actual extent of memory is an order of magnitude longer than predicted by Hsp90-inspired HMMs (e.g.  $t_{mem} \approx 40$  s (actual) versus  $t_{mem} \approx 6$  s (HMM) at  $\tau = 1$  s) and also displays (at least) two distinct time-scales (note the shoulder in  $\mathcal{D}_{\tau,q_0}(t)$  shown in figures 5). Moreover, the HMM falls short in accurately describing the correlation time ( $t_{corr} \approx 90$  s (actual) versus  $t_{\rm corr} \approx 14.5 \, {\rm s}({\rm HMM}))$ , and fails to capture the observed clear distinction between the open (orange) and closed (blue) conformation. The choice of the HMM thus directly affects the memory profile. A model that fails to capture the memory profile of the observed dynamics an intrinsic signature of the *actual* coupling of the observable to hidden degrees of freedom does not provide an adequate representation of the underlying microscopic dynamics.

We note that as a result of limited resolution and sampling, in our analysis we decided to bin the data in two, four, and eight bins respectively, whereby the binning had no significant effect on the memory profile past the expected fact that smaller binning introduced larger uncertainty (see green curves in figure 7). We stress that binning is *not* a requirement of the approach (i.e. one can use an arbitrarily fine binning). It is the limited sampling that causes the need for binning. This may be important to consider, since binning itself may introduce memory [13, 24]. However, the effect of binning alone is conceptually easily assessed. If it plays a significant role, the inferred memory profile will show a strong dependence on the size of the bins.

In the case of Hsp90 we expect very similar memory profiles also for 'point-like' binning (i.e. if one infers the respective probability densities with continuous support). For an experimental example where this is very clearly demonstrated see analysis of force-spectroscopy data on DNA-hairpin folding in [18]. The memory persists at arbitrarily fine binning because making bins finer only reduces lumping of the plasmon signal that, however, itself inherently corresponds to many microscopic states lumped into the same plasmon signal.

Our analysis thus suggests that the transition between opening and closing of Hsp90 occurs on timescales influenced by memory, and therefore is affected by the initial conditions of the hidden degrees of freedom.

#### 4. Discussion and conclusions

The conformational dynamics of proteins are often measured between two sites in a protein by single-molecule methods. The resulting time-series are then typically analysed by Hidden-Markov based methods, without properly assessing the representation of the coupling of the observable with hidden degrees of freedom that is parametrically encoded in the underlying full Markov model. Here, we analysed long single-molecule time-series of the conformational dynamics of the Hsp90 protein probed by a plasmon ruler. In particular, we carried out a model-free analysis of the non-Markovian behaviour and conclusively confirmed and quantified the memory in the observed dynamics. This was presumably the first time that the conformational dynamics of a protein were quantified in a model-free manner and in such detail in an experimental setting (model-based approaches have been used before [49-53]). Moreover, our analysis allowed us to quantify the lifetime of the memory in the time-series of Hsp90. Strikingly, the memory was confirmed to be present for up to 50s. Recall that the effect of the coupling of the gold nanoparticles to Hsp90's dynamics on the investigated time scales could already be excluded in the original publication [3], such that we can attribute the memory to Hsp90 dynamics alone. The here quantified duration of memory is on a time scale which is significant for many biological processes of the observed protein, therefore it is conceivable (if not very likely) that memory has biological relevance and implications. This means, for example, that interactions with client proteins are affected by Hsp90's conformational dynamics and vice versa. Thus, Hsp90's interactions have long lasting effects and are influenced by its dynamics and interactions for up to 50s. To set the timescale into perspective, human RNA is translated to a protein at a rate of approximately 5 amino acids per second [54]. For a protein consisting of 400 amino acids, translation would take about 80s. If Hsp90 — fulfilling its chaperone function — assists in folding this protein into its mature form, the starting conformation of Hsp90 would directly affect how Hsp90 assists in the folding process. Furthermore, our analysis is also suitable for inferring memory in non-equilibrium (driven) systems as we have exemplified in figure 3. We anticipate that this will be very important in the future, as compared to Hsp90's ATPase rate (1 ATP per min [55, 56]) memory of 50s is still long with likely significant impact on Hsp90's non-equilibrium dynamics.

We note that the most general, non-linear versions of Zwanzig-Mori projection-operator methods with *non-linear drifts and explicitly time-dependent and non-linear (i.e. non-convolution-type) memory kernels* (see e.g. equation (33) in [57], equation (9.58) in [58],

equation (2.7) in [59] or equations (6-7) in [60]) may formally also be considered as model free as they are essentially identities of the Liouville equation. However, the exact equations of motions obtained in an approximation-free manner are nominally *not* closed, and a model for the non-equilibrium initial conditions of hidden degrees is required for closure and averaging. In practice, space-independent and linear convolution-type memory kernels are assumed, which may thus not be faithfully considered as truly being model free. Furthermore, linear generalized Langevin equations for the auto-correlation C(t) may formally also be considered as model free [61]. However, extending the explanation in the Introduction, the resulting memory kernel generally does *not* reflect memory, only the presence of multiple time-scales in the attenuation of correlations (for a direct counterexample see figure 4(a) in [19]). One may, of course, nevertheless infer a 'memory time' from the respective memory kernels, which, however, has a different meaning than the memory time we investigate here. The relations between the respective memory times remain an interesting open question for future work.

Altogether, our study shows that for the analysis of single-molecule time-series, e.g. of protein dynamics, it is crucial to initially investigate whether memory is present. Only in the absence of memory on relevant timescales, Markov modelling is justified, because it relies on the assumption of a memoryless process. Additionally, (hidden) Markov modelling relies on assumptions about models and—in the context of the model selection process and all its associated challenges [23, 52, 62] — one typically selects the one with the highest likelihood, potentially introducing bias and explicitly setting the memory profile. A HMM that fails to correctly reproduce the memory profile (i.e. its magnitude and especially duration) cannot be considered as an adequate representation of the underlying dynamics, in particular the coupling of the observable to hidden degrees of freedom, and thus cannot provide reliable insight. Modelfree methods as described here, on the other hand, can overcome this bias given sufficient statistics. Transition-path based approaches [63] could then be used. Our data and analysis indicate a pressing need for such model-free methods that have to be further developed and optimised to analyse single-molecule time-series. On the positive side, our work demonstrates that such an analysis, despite requiring substantial statistics, is already within experimental reach.

#### Data availability statement

The data that support the findings of this study are openly available at the following URL/DOI: https://zenodo.org/records/12697153 [42].

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# Appendix A. Fit parameters for correlation time $t_{corr}$ and memory time $t_{mem}$

**Table 1.** Correlation timescale  $t_{corr}$  and memory timescale  $t_{mem}$  for the three different simulated model systems (here  $q_0 = 0.25$  and different  $\tau$ ). Values are obtained by fitting the long-time behaviour of C(t) and  $\mathcal{D}_{\tau,q_0}(t)$ , respectively, to the exponential  $\simeq \exp(-t/t_i)$  where  $i \in \{\text{corr, mem}\}$ .

System	<i>t</i> <sub>corr</sub> [s]	$t_{\rm mem}$ [s] $ au = 1.0  { m s}$	$t_{\rm mem}$ [s] $ au = 2.0  { m s}$	$t_{\rm mem}$ [s] $ au = 5.0  { m s}$
Model 1	$     \begin{array}{r}             49.11 \pm \\             5.94 \times 10^{-5}         \end{array}     $	_		
		$t_{ m mem}$ [s] au = 0.5   m s	$t_{\rm mem}$ [s] au = 1.0 s	$t_{\rm mem}$ [s] au = 2.5 s
Model 2	$\begin{array}{c} 14.53 \pm \\ 3.50 \times 10^{-5} \end{array}$	$6.29 \pm 4.89 \times 10^{-2}$	$6.50 \pm 2.05 \times 10^{-2}$	$6.20 \pm 9.50 \times 10^{-1}$
		$t_{\rm mem}$ [s] $ au = 0.2  { m s}$	$t_{\rm mem}$ [s] $ au = 0.4  { m s}$	$t_{\rm mem}$ [s] au = 1.0 s
Model 3	$\begin{array}{c} 3.41 \pm \\ 2.78 \times 10^{-4} \end{array}$	$2.89 \pm 3.37 \times 10^{-2}$	$2.14 \pm 3.64 \times 10^{-3}$	$2.09 \pm 3.45 \times 10^{-4}$

**Table 2.** Correlation timescale  $t_{\text{corr}}$  and memory timescale  $t_{\text{mem}}$  for the experimental time-series of Hsp90 (here  $q_0 = 0.25$  or  $q_0 = 0.75$ , and different  $\tau$ ). Values are obtained by fitting the long-time behaviour of C(t) and  $\mathcal{D}_{\tau,q_0}(t)$ , respectively, to the exponential  $\simeq \exp(-t/t_i)$  where  $i \in \{\text{corr, mem}\}$ .

System	<i>t</i> <sub>corr</sub> [s]	$t_{\rm mem}$ [s] $\tau = 1.0$ s	$t_{\rm mem}$ [s] $\tau = 3.0  {\rm s}$	$t_{\rm mem}$ [s] $\tau = 5.0  {\rm s}$
$\overline{\text{Hsp90}(q_0 = 0.25)}$	$\overline{90.00\pm0.42}$	$\overline{37.22\pm14.69}$	$\overline{45.98\pm12.81}$	$44.74 \pm 17.20$
		$t_{\rm mem}$ [s] $\tau = 1.0$ s	$t_{\rm mem}$ [s] $\tau = 3.0  {\rm s}$	$t_{\rm mem}$ [s] $\tau = 5.0  {\rm s}$
Hsp90 $(q_0 = 0.75)$		$\overline{39.18\pm16.71}$	$\overline{47.92\pm13.65}$	$45.71 \pm 18.46$



Appendix B. Memory analysis for simulated systems for  $q_0 = 0.75$ 

**Figure 6.** Analysis of memory in projected dynamics for three different simulated model systems with 100 time-series of 6 h duration each. The Kullback–Leibler divergence  $\mathcal{D}_{\tau,q_0}(t)$  in equation (6) between the transition probability density of the observed dynamics  $G(q,t|q_0)$  and its Chapman–Kolmogorov construction  $G_{\tau}^{CK}(q,t|q_0)$  as a function of time *t* for underlying dynamics of (a)–(c) a four-state Markov model with uniform transitions, (d)–(f) a four-state Markov model with Hsp90-inspired transition dynamics that obey detailed balance, and (g)–(i) a four-state Markov model with Hsp90-inspired transition dynamics set here to  $q_0 = 0.75$ , corresponding to the closed conformation of Hsp90 (see main text for  $q_0 = 0.25$ ). Memory in the Chapman–Kolmogorov construct is reset to zero at different times  $\tau$  and the memory time-scale  $t_{mem}$  (red marker) is obtained by fitting the long-time behaviour according to  $\mathcal{D}_{\tau,q_0}(t) \simeq \exp(-t/t_{mem})$ . Shaded areas depict standard deviations  $\sigma_{\mathcal{D}}(t)$ , obtained by averaging over 100 independent simulated trajectories and insets depict the same analysis for only one trajectory that is 100 times as long.



# Appendix C. Memory analysis for experimental time-traces with different spatial coarse-graining

**Figure 7.** Analysis of memory in experimental Hsp90 time-series. Kullback–Leibler divergence  $\mathcal{D}_{\tau,q_0}(t)$  in equation (6) between the transition probability density of the observed dynamics  $G(q,t|q_0)$  and its Chapman-Kolmogorov construction  $G_{\tau}^{CK}(q,t|q_0)$  as a function of time *t*. Four different initial conditions  $q_0$  are considered, corresponding to starting in the left (a)–(c) and right (d)–(f) peak in the probability density  $p_{eq}(q)$ , respectively. The four bins (blue, orange) are of equal width  $l_{bin} = 0.25$  and centred at 0.125, 0.375, 0.625, and 0.875. Memory in the Chapman-Kolmogorov construction is reset to zero at times (a), (d)  $\tau = 1$  s, (b), (e)  $\tau = 3$  s, and (c), (f)  $\tau = 6$  s, respectively. Error bars depict the standard deviation  $\sigma_{\mathcal{D}}(t)$ , obtained by systematically neglecting 20% of the recorded data for M = 100 independent bootstrapped repetitions and subsequent averaging. Black lines depict  $\mathcal{D}_{\tau,q_0}(t)$  as shown in figure 5 and corresponds to the initial condition (a)–(c)  $q_0 = 0.75$  and (d)–(f)  $q_0 = 0.25$ , respectively. Green lines show  $\mathcal{D}_{\tau,q_0}(t)$  for spatial lumping into eight equal bins of width  $l_{bin} = 0.125$  and  $\tau = 1$  s. Due to limited sampling the analysis was inconclusive for  $\tau = 3$  s and  $\tau = 6$  s.

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